

STATISTICAL INFERENCE (MA862)

Lecture Slides

Topic 1: Monte Carlo Simulation

Example

- Let us want to estimate the average distance between two randomly selected points in a region.
- Let $\mathbf{X} = (X_1, X_2)$ and $\mathbf{Y} = (Y_1, Y_2)$ be independent and uniformly distributed two points drawn from a finite rectangle $R = [0, a] \times [0, b]$.
- The Euclidean distance between these two points is

$$Z = d(\mathbf{X}, \mathbf{Y}) = \sqrt{(X_1 - Y_1)^2 + (X_2 - Y_2)^2}.$$

- We need $E(Z)$.

$$E(Z) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \sqrt{(x_1 - y_1)^2 + (x_2 - y_2)^2} \\ \times f_{X_1, X_2}(x_1, x_2) f_{Y_1, Y_2}(y_1, y_2) dx_1 dx_2 dy_1 dy_2$$

Example

- We can approximate $E(Z)$ by sampling pair points $(\mathbf{X}_i, \mathbf{Y}_i)$, $i = 1, 2, \dots, n$, from R and then calculating the average

$$\frac{1}{n} \sum_{i=1}^n d(\mathbf{X}_i, \mathbf{Y}_i).$$

Simple Monte Carlo

- In a simple Monte Carlo problem, we express the quantity we want to know as the expected value of a random variable Y , such as $\mu = E(Y)$.
- Generate values Y_1, \dots, Y_n independently from the distribution of Y .
- Take their average

$$\hat{\mu}_n = \frac{1}{n} \sum_{i=1}^n Y_i$$

as an estimate of μ .

Simple Monte Carlo

- In many examples, $Y = h(\mathbf{X})$.
- \mathbf{X} has a PMF/PDF $p(\mathbf{x})$ (known).
- f is a real-valued function defined over the support of \mathbf{X} .

Justification of Simple MC

- The strong law of large numbers:

$$\mathbb{P} \left(\lim_{n \rightarrow \infty} |\hat{\mu}_n - \mu| = 0 \right) = 1.$$

- Loosely speaking, the SLLNs says that the error in approximation will be very small if we increase n .

Random Number Generation

- Our aim is to generate random number from appropriate distribution.
- Basic step of a random number generation from a distribution is to generate random numbers from Uniform distribution $U(0, 1)$.
- The PDF of a random variable having $U(0, 1)$ distribution is

$$f(x) = \begin{cases} 1 & \text{if } 0 < x < 1 \\ 0 & \text{otherwise.} \end{cases}$$

Uniform Random Numbers Generation

- To generate random numbers from a process that according to well established understanding of physics is truly random.
- Radioactive particle emission, that are thought to be truly random.
- Such process has it's own draw back.
- Therefore, people use Pseudo-random numbers (computer generated).
- We will not discuss it in detail, as almost all software has a routine to generate pseudo-random numbers from $U(0, 1)$.
- Now-onward, the pseudo-random number will be referred as random number.

Uniform Random Numbers Generation

- A linear congruence generator (LCG):

$$x_{i+1} = ax_i \bmod m, \quad u_{i+1} = \frac{x_{i+1}}{m}, \quad i = 0, 1, 2, \dots$$

- The multiplier a and the modulus m are integer constants.
- The initial value (*seed*) x_0 is called **seed**.
- x_0 is an integer between 1 and $m - 1$.
- LCG is a deterministic recurrence relation.

Uniform Random Numbers Generation

- The general linear congruence generator (GLCG):

$$x_{i+1} = (ax_i + c) \bmod m, \quad u_{i+1} = \frac{x_{i+1}}{m}, \quad i = 0, 1, 2, \dots$$

- a , m and c are appropriate integers.

Non-uniform Random Number Generation

- Method for transforming random numbers from $U(0, 1)$ distribution to samples from other required distributions.
- The two most widely used general techniques are:
 - ① Inverse Transform Method.
 - ② Acceptance Rejection Method.

Inverse Transform Method

- The inverse transform method is based on the following theorem.

Theorem 1: Let F be a CDF. Define the quasi-inverse of F by

$$F^{-1}(u) = \inf \{x \in \mathbb{R} : F(x) \geq u\} \quad \text{for } 0 < u < 1.$$

Let $U \sim U(0, 1)$ and $X = F^{-1}(U)$. Then, the CDF of X is F .

- $\{x \in \mathbb{R} : F(x) \geq u\}$ is non-empty and has a lower bound for all $u \in (0, 1)$.

Inverse Transform Method

- Want a sample from the CDF $F(x)$. That means that we want to generate a random variable X with the property that $P(X \leq x) = F(x)$ for all $x \in \mathbb{R}$.
- Using above theorem, we have the following algorithm.

Algorithm 1 Inverse Transform Method

- 1: Generate U from $U(0, 1)$ distribution.
 - 2: Set $X = F^{-1}(U)$.
 - 3: Return X .
-

- In principle, this algorithm can be used to generate random number from any distribution.
- However, there are computational aspects. We generally use this algorithm if F^{-1} is in closed form and easy to compute.

Inverse Transform Method

Example 1: (Generation from exponential distribution) The PDF is

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x > 0 \\ 0 & \text{otherwise.} \end{cases}$$

Example 2: (Generation from Arc Sin Law) Consider the CDF

$$F(x) = \frac{2}{\pi} \arcsin \sqrt{x}, \quad 0 \leq x \leq 1.$$

Example 3: (Generation from Rayleigh Distribution) The CDF is

$$F(x) = 1 - e^{-2x(x-b)}, \quad x \geq b.$$

Inverse Transform Method

Lemma 1: F and F^{-1} both are non-decreasing.

Lemma 2: $F F^{-1}(u) \geq u$ for all $u \in (0, 1)$.

Lemma 3: $F^{-1} F(x) \leq x$ for all $x \in \mathbb{R}$.

Lemma 4: For $x \in \mathbb{R}$ and $0 < u < 1$, $F(x) \geq u$ if and only if $F^{-1}(u) \leq x$.

Inverse Transform Method

Example 4: Generation from a Bernoulli distribution with probability of success p ($q = 1 - p$).

$$F_X(x) = \begin{cases} 0 & \text{if } x < 0 \\ q & \text{if } 0 \leq x < 1 \\ 1 & \text{if } x \geq 1. \end{cases} \quad \text{and} \quad F^{-1}(u) = \begin{cases} 0 & \text{if } 0 < u < q \\ 1 & \text{if } q \leq u < 1. \end{cases}$$

Algorithm 2 Generation from *Bernoulli*(p)

- 1: generate U from $U(0, 1)$.
- 2: **if** $U < 1 - p$ **then**
- 3: Set $X \leftarrow 0$.
- 4: **else**
- 5: Set $X \leftarrow 1$.
- 6: **end if**
- 7: **return** X

Inverse Transform Method

Example 5: Generation from a discrete distribution with finite support.

- ▶ Consider a DRV X whose support is $c_1 < c_2 < c_3 < \dots < c_N$.
- ▶ Let $p_i = P(X = c_i)$, $i = 1, 2, 3, \dots, N$.
- ▶ Set $q_0 = 0$ and $q_i = \sum_{j=1}^i p_j$, $i = 1, 2, 3, \dots, N$.

Algorithm 3 Inversion Transformation Method for Discrete Random Variable with Finite Support

- 1: Generate a uniform $U \sim U(0, 1)$.
 - 2: Find $K \in \{1, 2, \dots, N\}$ such that $q_{K-1} < U \leq q_K$.
 - 3: Return c_K .
-

Acceptance-Rejection Method

- We want to generate random number from a PDF f (target distribution).
- Let g (candidate distribution) be a PDF such that for all $x \in \mathbb{R}$ and for some $c \geq 1$

$$f(x) \leq cg(x).$$

- The technique to generate random number from g is known.
- Then we can use the following algorithm to generate random number from f .

Algorithm 4 Acceptance Rejection Method

- 1: **repeat**
- 2: generate X from distribution g .
- 3: generate U from $U(0, 1)$.
- 4: **until** $U \leq \frac{f(X)}{cg(X)}$
- 5: **return** X

Generation from Gamma Distribution

Example 6: Generation random number from $\text{Gamma}(\alpha, \beta)$.

► The PDF of the distribution is

$$f(x) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x} \quad \text{for } x > 0.$$

● We will consider generation from $\text{Gamma}(\alpha, 1)$ for $\alpha > 0$.

Generation from Gamma Distribution

Case I : $0 < \alpha < 1$.

► Take

$$g(x) = \begin{cases} \frac{x^{\alpha-1}}{A} & \text{if } 0 < x < 1 \\ \frac{e^{-x}}{A} & \text{if } x \geq 1, \end{cases}$$

where $A = \frac{1}{\alpha} + \frac{1}{e}$.

► Then, $f(x) \leq cg(x)$, where $c = \frac{A}{\Gamma(\alpha)}$.

► The CDF corresponding to g is

$$G(x) = \begin{cases} \frac{x^\alpha}{\alpha A} & \text{if } 0 < x < 1 \\ 1 - \frac{e^{-x}}{A} & \text{if } x \geq 1. \end{cases}$$

► Now,

$$G^{-1}(u) = \begin{cases} (\alpha Au)^{\frac{1}{\alpha}} & \text{if } 0 < u < \frac{1}{\alpha A} \\ -\ln A - \ln(1 - u) & \text{if } \frac{1}{\alpha A} \leq u < 1. \end{cases}$$

Generation from Gamma Distribution

Algorithm 5 Generation from $\text{Gamma}(\alpha, 1)$ for $0 < \alpha < 1$

```
1: repeat
2:   generate  $U_1$  from  $U(0, 1)$ 
3:   if  $U < \frac{1}{\alpha A}$  then
4:     Set  $X \leftarrow (\alpha AU)^{\frac{1}{\alpha}}$ 
5:   else
6:     Set  $X \leftarrow -\ln A - \ln(1 - U)$ 
7:   end if
8:   generate  $U_2$  from  $U(0, 1)$ 
9: until  $cg(X)U_2 \leq f(X)$ 
10: return  $X$ 
```

Generation from Gamma Distribution

Case II: α is a positive integer.

▶ Let $X_i \stackrel{i.i.d.}{\sim} \text{Exp}(1)$ for $i = 1, 2, \dots, n$.

▶ Then $\sum_{i=1}^n X_i \sim \text{Gamma}(n, 1)$.

Algorithm 6 Generation from $\text{Gamma}(\alpha, 1)$, α is a positive integer

- 1: Set $n \leftarrow \alpha$ and $Y \leftarrow 0$
- 2: **while** $n \neq 0$ **do**
- 3: generate U from $U(0, 1)$
- 4: Set $X \leftarrow -\ln(U)$
- 5: $Y \leftarrow Y + X$
- 6: $n \leftarrow n - 1$
- 7: **end while**
- 8: **return** Y

Generation from Gamma Distribution

Case III: $\alpha > 1$ and not an integer.

- ▶ Let $X \sim \text{Gamma}(\alpha_1, \beta)$ and $Y \sim \text{Gamma}(\alpha_2, \beta)$.
- ▶ Suppose that X and Y are independent.
- ▶ Then $X + Y \sim \text{Gamma}(\alpha_1 + \alpha_2, \beta)$.
- ▶ $\lfloor x \rfloor$: the integer part of the positive real number x .
- ▶ $\{x\}$: the fractional part of the positive real number x .
- ▶ $\alpha = \lfloor \alpha \rfloor + \{\alpha\}$.

Algorithm 7 Generation from $\text{Gamma}(\alpha, 1)$ when $\alpha > 1$ and α is not an integer

- 1: generate Y from $\text{Gamma}(\{\alpha\}, 1)$ using Algorithm 5
 - 2: generate X from $\text{Gamma}(\lfloor \alpha \rfloor, 1)$ using Algorithm 6
 - 3: $Z = X + Y$
 - 4: **return** Z
-

Acceptance-Rejection Method

Theorem 2: Let f and g be two PDFs such that

$$f(x) \leq cg(x) \quad \text{for all } x \in \mathbb{R} \text{ and for some } c \geq 1.$$

Then X generated by Algorithm 4 has PDF f .

Technique based on Transformation

Example 7: Generation of random number from $Beta(\alpha, \beta)$.

- ▶ $X \sim Gamma(\alpha_1, \beta)$
- ▶ $Y \sim Gamma(\alpha_2, \beta)$
- ▶ X and Y are independent
- ▶ Then $\frac{X}{X + Y} \sim Beta(\alpha_1, \alpha_2)$.

Algorithm 8 Generation from $Beta(\alpha_1, \alpha_2)$ distribution

- 1: generate X from $Gamma(\alpha_1, \beta)$
 - 2: generate Y from $Gamma(\alpha_2, \beta)$
 - 3: $Z = \frac{X}{X + Y}$
 - 4: **return** Z
-

Technique based on Transformation

Example 8: Generation of random number from $N(\mu, \sigma^2)$

▶ $U_1, U_2 \stackrel{i.i.d.}{\sim} U(0, 1)$

▶ Define

$$Z_1 = \sqrt{-2 \ln U_1} \cos(2\pi U_2) \quad \text{and} \quad Z_2 = \sqrt{-2 \ln U_1} \sin(2\pi U_2).$$

▶ Then $Z_1, Z_2 \stackrel{i.i.d.}{\sim} N(0, 1)$

▶ This transformation is called **Box-Muller transformation**.

Algorithm 9 Box-Muller Method to generate from $N(0, 1)$

1: generate U_1 and U_2 from $U(0, 1)$

2: $R \leftarrow \sqrt{-2 \ln U_1}$

3: $\theta \leftarrow 2\pi U_2$

4: $Z_1 \leftarrow R \cos(\theta)$

5: $Z_2 \leftarrow R \sin(\theta)$

6: **return** (Z_1, Z_2) .

Technique based on Transformation

Example 9: Generation from $Geometric(p)$.

- ▶ The PMF is given by

$$P(X = i) = p(1 - p)^i \quad \text{for } i = 0, 1, 2, \dots$$

- ▶ Let Y be an exponential random variable with mean $\frac{1}{\lambda}$.
- ▶ Take $W = \lfloor Y \rfloor$. Then

$$P(W = i) = e^{-i\lambda} (1 - e^{-\lambda}) \quad \text{for } i = 0, 1, 2, \dots$$

- ▶ $W \sim Geometric(1 - e^{-\lambda})$.

Algorithm 10 Generation from $Geometric(p)$

- 1: generate U from $U(0, 1)$
 - 2: $X \leftarrow \left\lfloor \frac{\ln U}{\ln(1-p)} \right\rfloor$
 - 3: **return** X .
-

Error in Monte Carlo Integration

- Both laws of large numbers tell us that Monte Carlo will eventually produce an error as small as we like.
- They do not tell us how large n has to be for this to happen.
- They also do not say for a given sample Y_1, \dots, Y_n whether the error is likely to be small.
- The situation improves markedly when Y has a finite variance.
- Suppose that $\text{Var}(Y) = \sigma^2 < \infty$.
- The mean of $\hat{\mu}_n$ is $\mathbb{E}(\hat{\mu}_n) = \frac{1}{n} \sum_{i=1}^n \mathbb{E}(Y_i) = \mu$.
- The variance of $\hat{\mu}_n$ is $\text{Var}(\hat{\mu}_n) = \mathbb{E}((\hat{\mu}_n - \mu)^2) = \frac{\sigma^2}{n}$.
- The variance of $\hat{\mu}$ may be considered as (average) error in $\hat{\mu}$ for approximating or estimating μ .
- For fixed n , the variance increases as σ increases.
- The variance decreases as n increases for fixed σ .
- The root mean squared error of $\hat{\mu}_n$ is $\sqrt{\text{Var}(\hat{\mu}_n)} = \frac{\sigma}{\sqrt{n}}$.